

# On the fake stationary rough volatility models.

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The *short/rough memory of Volterra Integral Diffusions*  
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In the Volterra Heston model, the asset price process  $S$  and the variance process  $V$  are given by

$$\frac{dS_t}{S_t} = \sqrt{V_t} (\rho dW_t + \sqrt{1 - \rho^2} dW_t^\top), \quad S_0 \in (0, \infty), \quad (1)$$

$$V_t = \sigma^2(X_t), \quad X_t = X_0 + \int_0^t K_\alpha(t-s) \left( (\mu(s) - \lambda X_s) ds + \nu \sqrt{V_s} dW_s \right). \quad (2)$$

where  $K_\alpha$  is the fractional kernel  $\alpha := H + \frac{1}{2}$ ,  $W = (W, W^\top)$  is a two-dimensional independent standard Brownian motion  $\rho \in [-1, 1]$ ,  $\mu$  a deterministic function,  $\lambda, \nu \in \mathbb{R}_+$ .

- If  $\sigma(x) = \sqrt{x}$  and  $\alpha \in (\frac{1}{2}, 1)$ , we recover the standard **rough heston model**
- If  $\sigma(x) = \sqrt{\kappa_0 + \kappa_2(x-a)^2}$  and  $\alpha \in (\frac{1}{2}, 1)$ , we recover the **quadratic rough heston model**

## Potential challenges:

- Exhibits **distinct regimes** for short and long maturities
- Due to the **intrinsic non-stationarity** of the Volterra Equation (2), see, e.g. [Pagès, 2024]

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## Solution:

- Introduce the **“fake stationary Volterra” Heston model** in the **terminology** of [Pagès, 2024]
- Ensures a **time-consistent** modeling framework

# Fake stationary Volterra processes with affine drift and convolutive kernel

Fix  $T > 0$ , under the given **complete filtered probability space**  $(\Omega, \mathcal{F}, \mathbb{F} = \{\mathcal{F}_t\}_{0 \leq t \leq T}, \mathbb{P})$ , let  $X$  be the following  $\mathbb{R}$ -valued scaled **Volterra SDE** driven by the 1-dimensional Brownian motion  $W$ :

$$X_t = X_0 \varphi(t) + \int_0^t K(t-s)(\mu(s) - \lambda X_s) ds + \int_0^t K(t-s) \sigma(X_s) dW_s, \quad V_0 \perp\!\!\!\perp W. \quad (3)$$

Here,  $K \in L^2([0, T], \mathbb{R})$  is a scalar kernel,  $\varphi$  a **(locally) bounded Borel initial function** and  $\lambda \in \mathbb{R}_+$ ,  $\mu: \mathbb{R}_+ \rightarrow \mathbb{R}$ ,  $\sigma: \mathbb{R} \rightarrow \mathbb{R}$ .

**Assumption (Well-posedness):** We assume there exists at least a **continuous weak solution** to the Volterra SDE (3).

**Wiener–Hopf transform:** Let  $f_\lambda$  be the **resolvent of  $\lambda K$** , (i.e.  $f_\lambda + \lambda K * f_\lambda = \lambda K$ ), the solution  $(X_t)_{t \geq 0}$  of the Volterra SDE (3) also satisfies:

$$X_t = X_0(\phi(t) - \int_0^t f_\lambda(t-s)\phi(s) ds) + \frac{1}{\lambda} \int_0^t f_\lambda(t-s)\mu(s) ds + \frac{1}{\lambda} \int_0^t f_\lambda(t-s)\sigma(X_s) dW_s. \quad (4)$$

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## Definition 1 (Fake Stationarity)

The process  $(X_t)_{t \geq 0}$  with diffusion coefficient  $\sigma(X)$ , starting from  $X_0 \in L^2(\mathbb{P})$ , exhibit a **fake stationary regime of type I** if:

$$\forall t \geq 0, \quad \mathbb{E} X_t = \mathbb{E} X_0 =: m_0, \quad \text{Var}(X_t) = \text{Var}(X_0) =: v_0 \quad \text{and} \quad \mathbb{E} \sigma^2(X_t) = \bar{\sigma}_0^2. \quad (5)$$

## Theorem 2 (Autonomous diffusion coefficient $\sigma$ [Gnabeyeu and Pagès, 2025])

Assume  $X_0 \in L^2(\mathbb{P})$  and let  $(X_t)_{t \geq 0}$  be a solution to the scaled Volterra Equation (3) starting from  $X_0$  and satisfying

$$\forall t \geq 0, \quad \mathbb{E} X_t = m_0, \quad \text{Var}(X_t) = v_0 \geq 0 \quad \text{and} \quad \mathbb{E} \sigma^2(X_t) = \bar{\sigma}_0^2 > 0. \quad (6)$$

Set  $c := \frac{v_0}{\bar{\sigma}_0^2} > 0$  be such that  $c \geq \frac{1}{\lambda^2} \int_0^\infty f_\lambda^2(s) ds$ . Then, a necessary condition for the relations (6) to be satisfied is that

$$\forall t \geq 0, \quad \varphi(t) = 1 - \lambda \int_0^t K(t-s) \left( \frac{\mu(s)}{\lambda m_0} - 1 \right) ds. \quad (7)$$

and the *mean-reverting function*  $\mu$  in Equation (3) is uniquely determined and is given by

$$\mu(t) := \frac{d}{dt} \left( \int_0^t g_\lambda(t-s) r_\lambda(ds) \right) \quad \text{with} \quad g_\lambda(t) = \lambda m_0 \left( 1 - \sqrt{1 - \frac{1}{c \lambda^2} \int_0^t f_{\alpha, \lambda}^2(s) ds} \right) \quad \text{and} \quad (f_\lambda * r_\lambda)(t) = (r_\lambda * f_\lambda)(t) = 1 \quad (8)$$

provided the *functional resolvent of the first kind*  $r_\lambda$  of  $f_\lambda$  is well-defined, In such case, the process  $(X_t)_{t \geq 0}$  has constant mean  $m_0$  and a stationary variance  $v_0$  over time.

**Remark 1:** In such case, the Volterra Equation (3) simplify to

$$X_t = X_0 - (X_0 - m_0) \left( 1 - \sqrt{1 - \frac{1}{c\lambda^2} \int_0^t f_\lambda^2(s) ds} \right) + \frac{1}{\lambda} \int_0^t f_\lambda(t-s) \sigma(X_s) dW_s. \quad (9)$$

**Example 3** (The case of  $\alpha$ -fractional integration kernels  $K_\alpha(t) = \frac{t^{\alpha-1}}{\Gamma(\alpha)} \mathbf{1}_{\mathbb{R}_+}(t)$ )

Let  $\alpha \in (0, 1)$  and consider the fractional integration kernel  $K_\alpha(t) = \frac{t^{\alpha-1}}{\Gamma(\alpha)} \mathbf{1}_{\mathbb{R}_+}(t)$ , with  $f_{\alpha,\lambda}$  defined as the resolvent of  $\lambda K_\alpha$ . Let  $c \geq \frac{1}{\lambda^2} \int_0^\infty f_{\alpha,\lambda}^2(s) ds$ . In view of Equation (8), it follows that:

$$\mu(t) = g_{\alpha,\lambda}(t) + \frac{1}{\lambda} D^\alpha g_{\alpha,\lambda}(t) \quad \text{with} \quad g_{\alpha,\lambda}(t) = \lambda m_0 \left( 1 - \sqrt{1 - \frac{1}{c\lambda^2} \int_0^t f_{\alpha,\lambda}^2(s) ds} \right)$$

and  $D^\alpha := \frac{d}{dt} I^{1-\alpha}$  where  $D^\alpha$  and  $I^{1-\alpha}$  denote, respectively, the **Riemann–Liouville fractional derivative** of order  $\alpha$ , and the **Riemann–Liouville fractional integral** of order  $1 - \alpha$

**Remark 2:** One checks that  $\mu$  is continuous on  $(0, +\infty)$  and that  $\lim_{t \rightarrow +\infty} \mu(t) \in \mathbb{R}$ . Moreover, by **Hardy–Littlewood's Tauberian theorem for Laplace transform**, we have (heuristically),  $\mu(t) \underset{0+}{\sim} \frac{m_0 \Gamma(2\alpha)}{2c(2\alpha-1)\Gamma(\alpha)^3} t^{\alpha-1}$ .

**Potential challenges:**

- Highly restrictive, as it effectively dictates **explicit closed-form for both  $\phi$  and the mean-reversion term  $\mu$**  in the drift.

# A numerical illustration of intrinsic fake stationary

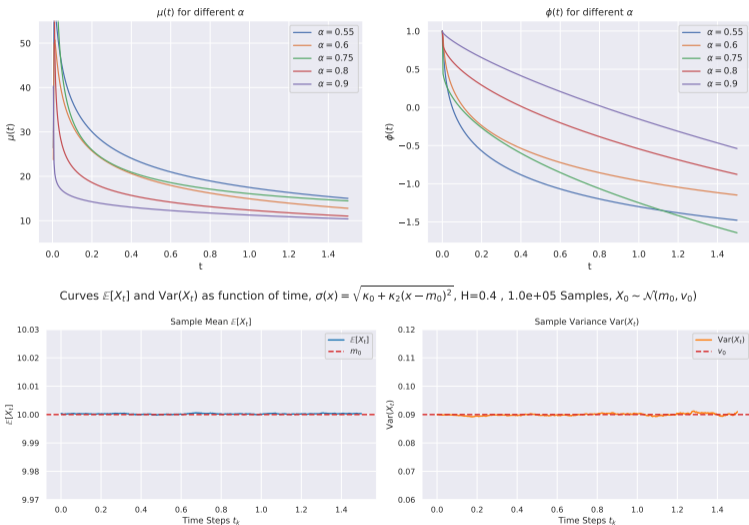


Figure: Top: Graphs of  $t \mapsto \mu(t)$  and  $t \mapsto \phi(t)$  for different values of  $\alpha$ . Bottom: Graphs of  $t_k \mapsto \mathbb{E}[X_{t_k}]$  and  $t_k \mapsto \text{Var}(X_{t_k}, M)$

## Fake stationary Volterra process using a stabilizer

$$X_t = X_0 \varphi(t) + \int_0^t K(t-s) \underbrace{(\mu(s) - \lambda X_s)}_{!!} ds + \int_0^t K(t-s) \varsigma(s) \sigma(X_s) dW_s, \quad V_0 \perp\!\!\!\perp W. \quad (10)$$

### Theorem 4 (Time-dependent diffusion coefficient [Gnabeyeu and Pagès, 2025].)

Let  $(X_t)_{t \geq 0}$  be a solution to the scaled Volterra equation in its form (3) starting from any random variable  $X_0 \in L^2(\Omega, \mathcal{F}, \mathbb{P})$ . Assume  $X_0 \in L^2(\mathbb{P})$  and let  $(X_t)_{t \geq 0}$  be a solution to (3) satisfying

$$\forall t \geq 0, \quad \mathbb{E} X_t = m_0 \neq 0, \quad \text{Var}(X_t) = v_0 \geq 0 \quad \text{and} \quad \bar{\sigma}(t) := \mathbb{E} \sigma^2(X_t) = \bar{\sigma}_0^2 > 0. \quad (11)$$

Then, a necessary and sufficient condition for the relations (11) to be satisfied is that

$$\forall t \geq 0, \quad \varphi(t) = 1 - \lambda \int_0^t K(t-s) \left( \frac{\mu(s)}{\lambda m_0} - 1 \right) ds. \quad (12)$$

and the couple  $(v_0, \varsigma(t))$  must satisfy the functional equation:

$$(E_{\lambda, c}): \quad \forall t \geq 0, \quad c \lambda^2 \left( 1 - \left( 1 - \frac{(f_\lambda * \mu)_t}{\lambda m_0} \right)^2 \right) = (f_\lambda^2 * \varsigma^2)(t) \quad \text{with} \quad c = \frac{v_0}{\bar{\sigma}_0^2} \quad \text{and} \quad \varsigma := \varsigma_{\lambda, c}. \quad (13)$$

### Definition 5 (Stabilizer)

We will call the **stabilizer (or corrector)** of the scaled stochastic Volterra equation (3) the (locally) bounded Borel function  $\varsigma := \varsigma_{\lambda, c}$ , solution (if any) to the **functional equation**  $(E_{\lambda, c})$  in (13).

$\varsigma_{\alpha,\rho,\lambda,c}^2$  for the  $\alpha$ -exponential fractional kernels  $K_{\alpha,\rho}(t) = e^{-\rho t} \frac{t^{\alpha-1}}{\Gamma(\alpha)} \mathbf{1}_{\mathbb{R}}(t)$ ,  $\rho \geq 0, \alpha \in (\frac{1}{2}, 1)$

- The computation of the function  $\varsigma_{\alpha,\rho,\lambda,c}^2$ , involves **knowing** the **form of the mean-reverting function  $\mu$**

We can **solve** the functional equation **numerically**

**Example 6 (Computing  $\varsigma_{\alpha,\rho,\lambda,c}^2$  with second-order extension of the triangular Volterra discretization)**

On a time grid  $t_k = k \frac{T}{n}$ ,  $k = 0, \dots, n$ , using a **piecewise linear interpolation** on  $[t_j, t_{j+1}]$ , we build the **recursive scheme**:  
 $\varsigma_{\alpha,\rho,\lambda,c}^2(0) = 0$  and  $\forall k \geq 1$ ,

$$\begin{aligned} c\lambda^2 \left( 1 - \left( 1 - \frac{(f_\lambda * \mu)(t_k)}{\lambda m_0} \right)^2 \right) &= (f_{\alpha,\rho,\lambda}^2 * \varsigma_{\alpha,\rho,\lambda,c}^2)(t_k) \\ &= \sum_{j=0}^{k-1} \left( \varsigma_{\alpha,\rho,\lambda,c}^2(t_j) \int_{t_j}^{t_{j+1}} f_{\alpha,\rho,\lambda}^2(t_k - s) \frac{t_{j+1} - s}{t_{j+1} - t_j} ds + \varsigma_{\alpha,\rho,\lambda,c}^2(t_{j+1}) \int_{t_j}^{t_{j+1}} f_{\alpha,\rho,\lambda}^2(t_k - s) \frac{s - t_j}{t_{j+1} - t_j} ds \right). \end{aligned} \quad (14)$$

which we can solve step by step (Lower-Triangular system) to **recover the values  $\varsigma_{\alpha,\rho,\lambda,c}^2(t_k)$ ,  $k \geq 1$** .

**Remark:** In this setting, the Volterra SDE (3) simplify to

$$X_t = X_0 - \frac{1}{\lambda m_0} (X_0 - m_0) \int_0^t f_\lambda(t-s) \mu(s) ds + \frac{1}{\lambda} \int_0^t f_\lambda(t-s) \varsigma_{\alpha,\rho,\lambda,c}(s) \sigma(X_s) dW_s. \quad (15)$$

# Illustrations with $\alpha$ -exponential fractional kernels $K_{\alpha,\rho}(t) = e^{-\rho t} \frac{t^{\alpha-1}}{\Gamma(\alpha)} \mathbf{1}_{\mathbb{R}}(t)$ , $\rho > 0, \alpha \in (\frac{1}{2}, 1)$

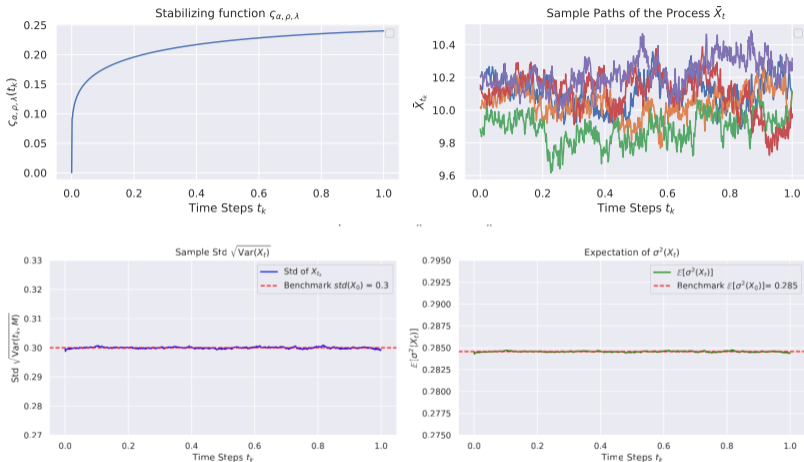


Figure: Stabilizer, sample paths, empirical variance, and mean squared volatility of  $(X_t)_{t \in [0,1]}$  with parameters:  $\sigma(x) = \sqrt{\kappa_0 + \kappa_2(x - m_0)^2}$ ,  $\mu(t) = \lambda m_0$ ,  $H = 0.1$ ,  $c = 0.316$ ,  $\lambda = 0.2$ ,  $\rho = 1.2$ ,  $m_0 = 10$ ,  $v_0 = 0.09$ ,  $\kappa_0 = 0.25$  and  $\kappa_2 = 0.384$ .

# Applications I: The fake stationary rough Heston model, $K(t) = K_\alpha(t) = \frac{\nu^{\alpha-1}}{\Gamma(\alpha)} \mathbf{1}_{\mathbb{R}}(t), \alpha \in (\frac{1}{2}, 1)$

In this setting, the pair  $(\log S, V)$ , where  $S$  denotes the asset price and  $V$  its variance process, is governed by

$$\begin{cases} \frac{dS_t}{S_t} = \sqrt{V_t} \left( \rho dW_t + \sqrt{1-\rho^2} dW_t^\top \right), & S_0 \in (0, \infty), \\ V_t = V_0 + \int_0^t K_\alpha(t-s) \left( \left(1 - \frac{V_0}{m_0}\right) \mu(s) - \lambda(V_s - V_0) \right) ds + \int_0^t K_\alpha(t-s) \nu \varsigma_{\alpha, \lambda, c}(s) \sqrt{V_s} dW_s. \end{cases} \quad (16)$$

with:

$$\forall t \geq 0, \quad c\lambda^2 \left( 1 - \left( 1 - \frac{(f_{\alpha, \lambda} * \mu)_t}{\lambda m_0} \right)^2 \right) = (f_{\alpha, \lambda}^2 * \varsigma_{\alpha, \lambda, c}^2)(t) \quad \text{with } c, \lambda > 0. \quad (17)$$

where,  $W = (W, W^\top)$  is a two-dimensional standard Brownian motion,  $\rho \in [-1, 1]$  a correlation,  $\mu$  a deterministic function,  $\lambda, \nu, c \in \mathbb{R}_+$  such that  $V$  is at least a weak solution to the Volterra equation (16) which writes equivalently:

$$V_t = V_0 + \frac{1}{\lambda} \int_0^t f_{\alpha, \lambda}(t-s) \left( 1 - \frac{V_0}{m_0} \right) \mu(s) ds + \frac{1}{\lambda} \int_0^t f_{\alpha, \lambda}(t-s) \nu \varsigma_{\alpha, \lambda, c}(s) \sqrt{V_s} dW_s, \quad \varsigma_{\alpha, \lambda, c}(t) \geq 0 \quad (18)$$

**Well-Posedness:** Unique-in-law positive continuous weak solution (as a scaling limit of a sequence of time-modulated Hawkes processes) see e.g., [Gnabeyeu et al., 2025].

### Assumption 7 (Integrability and uniform Hölder continuity)

Let  $\lambda > 0$  and  $\alpha \in (\frac{1}{2}, 1)$ . Assume the kernel  $K_\alpha$  is such that its resolvent  $f_{\alpha,\lambda}$  satisfies:

(i) Integrability:  $\int_0^{+\infty} f_{\alpha,\lambda}^{2\beta}(u) du < +\infty$  for some  $\beta > 1$ .

(ii) Hölder continuity:  $\exists \vartheta \in (0, 1]$  and  $C < +\infty$  such that  $\max_{i=1,2} \left[ \int_0^{+\infty} |f_{\alpha,\lambda}(u + \bar{\delta}) - f_{\alpha,\lambda}(u)|^i du \right]^{\frac{1}{i}} \leq C \bar{\delta}^\vartheta$ .

### Theorem 8 ([Gnabeyeu et al., 2026])

The fake stationary rough Heston model (16) has a **unique in law continuous**  $\mathbb{R}_+ \times \mathbb{R}_+$ -valued **weak solution**  $(S, V)$  for any initial condition  $(S_0, V_0) \in \mathbb{R}_+ \times L^2_{\mathbb{R}_+}(\mathbb{P})$  defined on some filtered probability space  $(\Omega, \mathcal{F}, (\mathcal{F})_{t \geq 0}, \mathbb{P})$ . Moreover, Let  $p > 0$  such that  $V_0 \in L^p(\mathbb{P})$ . Under Assumption 7, the sample paths of  $V$  are  $(\vartheta \wedge \frac{\beta-1}{2\beta} - \frac{1}{p} - \eta)$ -Hölder **pathwise continuous** for sufficiently small  $\eta > 0$  and for any  $T > 0$ , there exists a constant  $C_{p,T} > 0$  such that the following sup-norm holds

$$\left\| \sup_{t \in [0, T]} |V_t| \right\|_p^p = \mathbb{E} \left[ \sup_{t \in [0, T]} |V_t|^p \right] \leq C_{p,T} (1 + \mathbb{E}[|V_0|^p]). \quad (19)$$

Sketch of Proof: The **uniqueness in law** follows from the **characteristic function**. Since  $S$  is fully determined by  $V$ , the **existence of  $S$**  readily follows from that of  $V$ .

**Asymptotic Hawkes Setting:** Let  $(N_t^T, \Lambda_t^T)_{t \geq 0}$  be a sequence of 1-dimensional Hawkes processes indexed by  $T > 0$ , observed over  $[0, T]$ , and defined on the filtered probability space  $(\Omega^T, \mathcal{F}^T, \mathbb{P}^T, (\mathcal{F}_t^T)_{t \in [0, T]})$  with  $T \rightarrow \infty$  and the intensity process  $(\Lambda_t^T)_{t \geq 0}$  given by:

$$\Lambda_t^T = \underbrace{\Lambda_0^{*T} \Phi^T(t)}_{!!} + \mu^T(t) + \int_0^t \varphi^T(t-s) dN_s^T \quad (20)$$

where:

- $\mu^T(t)$  is the **non-negative baseline intensity**.
- $\varphi^T$  is a non-negative measurable and **completely monotonic** memory kernel with **regular varying tail** and  $\|\varphi^T\|_1 < \infty$ ,
- $\Phi^T(t) = \int_t^\infty \varphi^T(s) ds$  is the **non-negative initial condition**.
- $\Lambda_0^{*T}$  a **positive real-valued random variable**

## Ideas

Introduce a **time-dependent rescaling**. Define  $\Lambda_t^{*T} := \frac{1-a_T}{\bar{\mu}^T(\frac{t}{T})} \Lambda_{tT}^T$ , where  $(a_T)_{T \geq 0}$  is a  $(0, 1)$ -valued sequence of positive numbers converging to 1.

# Intuitions and main Heuristic convergence results

Let  $\Psi^T$  be the **resolvent** of  $\varphi^T$ . **Wiener–Hopf transform** of the intensity process  $(\Lambda_t^{*T})$  and **asymptotic limit as  $T \rightarrow +\infty$** .

$$\begin{aligned} \Lambda_t^{*T} = & \underbrace{\Lambda_0^{*T} (\Phi^T(tT) + (\Psi^T * \Phi^T)(tT)) \frac{1 - a_T}{\tilde{\mu}^T(\frac{t}{T})}}_{\xrightarrow{T \rightarrow +\infty} \Lambda_0^* (\phi(t) - \int_0^t f_{\alpha, \lambda}(t-s)\phi(s)ds)} + \frac{\mu^T}{\tilde{\mu}^T(\frac{t}{T})} \int_0^t \underbrace{T(1 - a_T)\Psi^T(T(t-s))}_{\xrightarrow{T \rightarrow +\infty} f_{\alpha, \lambda}(t-s)} \underbrace{\frac{\mu^T(Ts)}{\mu^T}}_{\xrightarrow{T \rightarrow +\infty} \frac{1}{\lambda}\mu(s)} ds \\ & + \underbrace{(1 - a_T) \frac{\mu^T(tT)}{\tilde{\mu}^T(\frac{t}{T})}}_{\xrightarrow{T \rightarrow +\infty} 0} + \int_0^t \underbrace{\sqrt{\frac{\tilde{\mu}^T(\frac{s}{T})}{T(1 - a_T)(\tilde{\mu}^T(\frac{t}{T}))^2}}}_{\xrightarrow{T \rightarrow +\infty} \frac{\nu}{\lambda}\zeta(s)} \underbrace{T(1 - a_T)\Psi^T(T(t-s))}_{\xrightarrow{T \rightarrow +\infty} f_{\alpha, \lambda}(t-s)} \sqrt{\Lambda_s^{*T}} dW_s^T, \forall t \in [0, t_0] \end{aligned}$$

with some constant  $\lambda, \nu \in \mathbb{R}_+^*$  and some **deterministic and continuous positive function**  $\phi, \mu$  and  $\zeta$ .

Taking the **limit as  $T \rightarrow \infty$** , we expect the limiting process  $\Lambda_t^*$  to be the solution (**with some additional assumptions**) of

$$\Lambda_t^* = \Lambda_0^* (\phi(t) - \int_0^t f_{\alpha, \lambda}(t-s)\phi(s)ds) + \frac{1}{\lambda} \int_0^t f_{\alpha, \lambda}(t-s)\mu(s)ds + \frac{1}{\lambda} \int_0^t f_{\alpha, \lambda}(t-s)\zeta(s)\sqrt{\Lambda_s^*}dW_s. \quad (21)$$

or equivalently (**Wiener–Hopf transform**),

$$\Lambda_t^* = \Lambda_0^* \phi(t) + \frac{1}{\Gamma(\alpha)} \int_0^t (t-s)^{\alpha-1} (\mu(s) - \lambda \Lambda_s^*) ds + \frac{\lambda \nu}{\Gamma(\alpha)} \int_0^t (t-s)^{\alpha-1} \zeta(s) \sqrt{\Lambda_s^*} dW_s. \quad (22)$$

**Formal Proof:** By **Functional Limits Theorem: "À la Jacod–Shiryaev"**, see [Gnabeyeu et al., 2026]

## Theorem 9 (Fourier–Laplace Transform of the fake stationary rough Heston model [Gnabeyeu et al., 2025])

Let  $\alpha \in (\frac{1}{2}, 1)$ ,  $\lambda, c > 0$  and  $u \in (\mathbb{C}^2)^*$ . Then there exists a unique function  $\psi \in L^2([0, T], \mathbb{C})$  solving the *fractional Riccati equation*

$$(D^\alpha \psi)(t) = \frac{1}{2}(-u^2 - iu) + (iu\rho\nu \varsigma_{\alpha, \lambda, c}(T-t) - \lambda) \psi(t) + \frac{\nu^2}{2} \varsigma_{\alpha, \lambda, c}^2(T-t) \psi^2(t), \quad t \in [0, T], \quad (I^{1-\alpha} \psi)(0) = 0. \quad (23)$$

leading to the Fourier–Laplace representation for the log-price:

$$\varphi_{T, V_0}(u) = \mathbb{E}_{V_0} \left[ e^{iu \log\left(\frac{S_T}{S_0}\right)} \right] = \exp \left( V_0 (I^{1-\alpha} \psi)(T) + \int_0^T \psi(T-s) \left( \left(1 - \frac{V_0}{m_0}\right) \mu(s) + \lambda V_0 \right) ds \right). \quad (24)$$

with  $D^\alpha = \frac{d}{dt} I^{1-\alpha}$  where  $D^\alpha$  and  $I^{1-\alpha}$  denote, respectively, the Riemann–Liouville *fractional derivative* of order  $\alpha$ , and the Riemann–Liouville *fractional integral* of order  $1 - \alpha$ .

### Practitioner corner:

- Generalized Adams–Bashforth–Moulton algorithm (fractional Adams method) for solving a fractional ODE (23).
- Plugging the numerical solution into (24) yields the characteristic function, from which standard Fourier methods allow the pricing of call and put options (Carr Madan, Lewis, etc.).

## Applications II: The fake stationary quadratic rough Heston (FS-QRH) model

**The Quadratic Rough Heston volatility Model**,  $K_\alpha(t) = \frac{u^{\alpha-1}}{\Gamma(\alpha)} \mathbf{1}_{\mathbb{R}}(t)$ ,  $\alpha \in (\frac{1}{2}, 1)$  and  $\alpha := H + \frac{1}{2}$

- has been introduced for the **joint calibration of the S&P 500 and VIX volatility surfaces**, accounting for the so-called Zumbach effect.
- typically used to model **options with short expiries**.

Let  $(X_t)_{t \geq 0}$  be a solution to the **fake stationary quadratic Volterra SDE**, where for every  $t > 0$ ,  $\underline{\mu}(t) = \lambda m_0$ , so that  $\varphi(t) = 1, \forall t \geq 0$ . We define the model

$$\begin{cases} \frac{dS_t}{S_t} = \sigma(X_t) (\rho dW_t + \sqrt{1 - \rho^2} dW_t^\top), \\ X_t = g_0(t) + \frac{1}{\lambda} \int_0^t f_{\alpha, \lambda}(t-s) \varsigma_{\alpha, \lambda, c}(s) \sigma(X_s) dW_s, \end{cases} \quad \text{with} \quad \begin{cases} \sigma(x) = \sqrt{\kappa_0 + \kappa_2(x - m_0)^2}, & \kappa_i \geq 0, i = 0, 2, \\ g_0(t) = m_0 + (X_0 - m_0) \left(1 - \int_0^t f_{\alpha, \lambda}(s) ds\right), \end{cases}$$

and  $\varsigma_{\alpha, \lambda, c}(t) \geq 0$ , solution to

$$(E_{\alpha, \lambda, c}): \quad \forall t \geq 0, \quad c\lambda^2 \left(1 - \left(1 - \int_0^t f_{\alpha, \lambda}(s) ds\right)^2\right) = (f_{\alpha, \lambda}^2 * \varsigma_{\alpha, \lambda, c}^2)(t) \quad \text{with} \quad \lambda > 0 \text{ and } c \in \left(0, \frac{1}{\kappa_2}\right). \quad (25)$$

$$X_0 \sim \mathcal{L}\left(m_0, v_0 = \frac{c\kappa_0}{1 - c\kappa_2}\right), \quad \underline{\text{Model Parameters:}} \quad \vec{\theta} := (\alpha, \lambda, \rho, c, m_0, \kappa_0, \kappa_2)$$

# Computing the Stabilizer $\varsigma_{\alpha,\lambda,c}$ for the $\alpha$ -fractional kernels $K_\alpha(t) = \frac{t^{\alpha-1}}{\Gamma(\alpha)} \mathbf{1}_{\mathbb{R}}(t)$ :

- $\Gamma(a) = \int_0^{+\infty} u^{a-1} e^{-u} du$ ,  $a > 0$  and  $B(a, b) = \int_0^1 u^{a-1} (1-u)^{b-1} du$ ,  $a, b > 0$ , the Gamma and the beta functions.

Set  $a_k = \frac{1}{\Gamma(\alpha k + 1)}$ ,  $b_k = \frac{1}{\Gamma(\alpha(k+1))}$ ,  $k \geq 0$ . The solution  $\varsigma_{\alpha,\lambda,c}^2$  to  $(E_{\alpha,\lambda,c})$  reads:

**Example 10 (Stabilizer  $\varsigma_{\alpha,\lambda,c}^2$ ,  $\alpha \in (\frac{1}{2}, 1)$ , see [Pagès, 2024])**

$$\varsigma_{\alpha,\lambda,c}^2(t) = c \lambda^{2-\frac{1}{\alpha}} \varsigma_\alpha^2(\lambda^{\frac{1}{\alpha}} t) \quad \text{where} \quad \varsigma_\alpha^2(t) := 2 t^{1-\alpha} \sum_{k \geq 0} (-1)^k c_k t^{\alpha k}. \quad (26)$$

with

$$c_0 = \frac{\Gamma(\alpha)^2}{\Gamma(2\alpha - 1)\Gamma(2 - \alpha)}, \quad \text{and for every } k \geq 1, c_k \text{ is defined inductively by:}$$

$$c_k = \frac{\Gamma(\alpha)^2 B(\alpha(k+1), 2(1-\alpha))}{\Gamma(2(1-\alpha))\Gamma(2\alpha - 1)} \left[ (a * b)_k - \alpha(k+1) \sum_{\ell=1}^k B(\alpha(\ell+2) - 1, \alpha(k-\ell-1) + 2) (b^{*2})_\ell c_{k-\ell} \right].$$

where for two sequences of real numbers  $(u_k)_{k \geq 0}$  and  $(v_k)_{k \geq 0}$ , the **Cauchy product** is defined as  $(u * v)_k = \sum_{\ell=0}^k u_\ell v_{k-\ell}$

Sketch of Proof: **Regular Variation ( Tauberian theorems)** on Laplace transforms of  $(E_{\alpha,\lambda,c})$ , see again [Pagès, 2024].

# Practitioner's corner: Truncating the power series defining $\varsigma_{\alpha,\lambda,c}^2$ at $n_{\max} = 50$

The left panel displays  $\log_{10} |c_k|$  as a function of  $k$  for  $\alpha \in [0.51, 0.99]$ . All curves fall below the float64 machine precision threshold, well before  $k = 50 \Rightarrow$  Truncating the series (26) at  $n_{\max} = 50$  is sufficient across the entire simulation range.

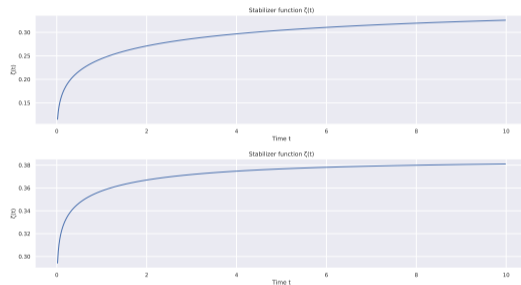
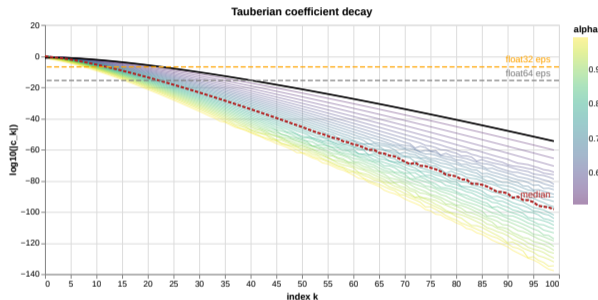


Figure: Left: Tauberian series coefficients over the time interval  $[0, 10]$  for Hurst exponents  $H \in (0, \frac{1}{2})$ . Right: Stabilizer mappings  $t_k \mapsto \varsigma_{\alpha,\lambda,c}(t_k)$  over the time interval  $[0, 10]$  for  $H = 0.1$  and  $H = 0.4$ . Parameters  $\lambda = c = 1$  and  $n = 600$ .

The right panel depicts the stabilizing functions  $t \mapsto \varsigma_{\alpha,\lambda,c}(t)$  for different values of  $\alpha \in (\frac{1}{2}, 1)$ .

## Applications II: Numerics for the fake stationary quadratic rough Heston model

On the **discrete grid**  $(t_k^n)_{0 \leq k \leq n}$   $t_k^n = \frac{kT}{n}$ ,  $k = 0, \dots, n$ ,

- for the Markovian log-stock price process  $\log S$ , we use the Euler–Maruyama scheme

$$\log S_{t_k^n} = \log S_{t_{k-1}^n} - \frac{1}{2} \frac{V_{t_{k-1}^n} + V_{t_k^n}}{2} \frac{T}{n} + \sqrt{V_{t_{k-1}^n}} \left( \rho \Delta W_{t_k^n} + \sqrt{1 - \rho^2} \Delta W_{t_k^n}^\perp \right), \quad (27)$$

where  $\Delta W_{t_k^n}^\perp \sim \mathcal{N}(0, \frac{T}{n})$  is independent of all other sources of randomness.

- for the Variance process, we use the **semi-integrated Euler scheme** which write recursively:

$$\bar{X}_{t_k^n}^n = g_0(t_k) + \frac{1}{\lambda} \sum_{\ell=0}^{k-1} \varsigma_{\alpha, \lambda, c}(t_{\ell+1}) \sigma(\bar{X}_{t_\ell^n}^n) \int_{t_\ell}^{t_{\ell+1}} f_{\alpha, \lambda}(t_k^n - s) dW_s \quad k = 1, \dots, n. \quad (28)$$

Let us denote by  $G = (G_{k\ell})_{k=0:n, \ell=0:n-1}$  the  $(n+1) \times n$  matrix involving the random terms  $I_k^{n, \ell} := \int_{t_\ell}^{t_{\ell+1}} f_{\alpha, \lambda}(t_k^n - s) dW_s$ .

$$(G_{k\ell})_{k=0:n, \ell=0:n-1} = \left( \begin{array}{c} \Delta W_{t_\ell} \mathbf{1}_{\{k=\ell\}}, k \leq \ell, \ell = 0 : n-1 \\ \int_{t_\ell}^{t_{\ell+1}} f_{\alpha, \lambda}(t_k - s) dW_s \mathbf{1}_{\{0 \leq \ell < k \leq n\}} \end{array} \right). \quad (29)$$

Then the following relation holds:  $\bar{X}_0^n = X_0$  and for every  $k = 1, \dots, n$

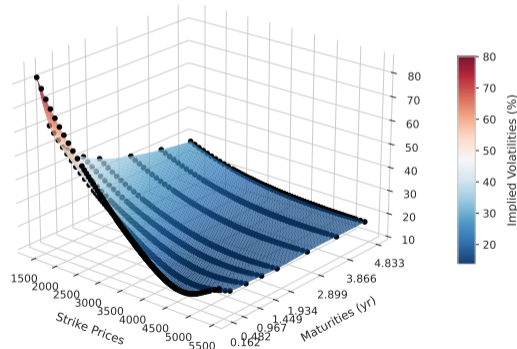
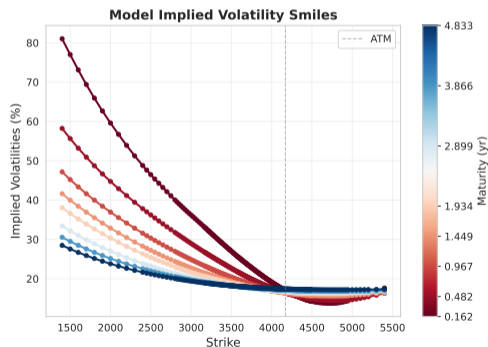
$$(\bar{X}_{t_k^n}^n)_{k=1:n} = (g_0(t_k))_{k=1:n} + G_{k, \cdot} (\varsigma_{\alpha, \lambda, c}(t_{\ell+1}) \sigma(\bar{X}_{t_\ell^n}^n))_{\ell=0:n-1} \quad (30)$$

by generating an independent sequence of **gaussian vectors**  $G_{k, \cdot}$ ,  $k = 0 \dots n$  using an **alternate extended and stable version of Cholesky decomposition** of a well-defined covariance matrix, see [Gnabeyeu and Pagès, 2025, Gnabeyeu and Pagès, 2026].

# Implied volatility surface generated by the FS-QRH model

Parameters	$\alpha = H + \frac{1}{2}$	$\lambda$	$\rho$	$c$	$m_0$	$\kappa_0$	$\kappa_2$
Fake stationary QRH	0.7478	1.4810	0.8575	0.2805	0.0752	0.0114	3.5652

Table: Parameters for the Fake Stationary Quadratic Rough Heston Model



- The model reproduces the patterns of different expiries, producing a **fairly good representation** of a market vol. surf.

## Fitting the FS-QRH model to the SPX volatility surface.

- We are left to calibrate our set of parameters  $\vec{\theta} := (\alpha, \lambda, \rho, c, m_0, \kappa_0, \kappa_2)$ : (A very small number of parameters).

⇒ Choice of the target objective function:<sup>1</sup> : Among Others, the **weighted squared-error loss**

$$F(\vec{\theta}) = \frac{1}{\#\sigma^{SPX}} \sum_{o \in \sigma^{SPX}} w_{vol}^o (\sigma^{o, mid} - \Psi_{\vec{\theta}}^{o, vol})^2, \quad \text{with} \quad w_{vol}^o := \frac{1}{\varepsilon + |\sigma^{o, bid} - \sigma^{o, ask}|}.$$

Here,  $\sigma^{SPX}$  the given set of SPX options,  $\sigma^{o, mid}$  the market mid implied volatility for the option  $o$ , and  $\Psi_{\vec{\theta}}^{o, vol}$  is the implied volatility of the option  $o$  in the FS-QRH model with the given parameter  $\vec{\theta}$  and obtained by **Monte-Carlo simulations**.

Our method of simulation ( the **semi-integrated Euler scheme** ) exhibit several advantages:

- its generality (even for long memory kernels),
- **accuracy** and **computational efficiency**

<sup>1</sup>Practitioners aim at **positioning the model implied volatility at the mid**.

# Model fit for very Short-Dated-Options (expiries measured in days)

Expiries	$\alpha = H + \frac{1}{2}$	$\lambda$	$\rho$	$c$	$m_0$	$\kappa_0$	$\kappa_2$
$T = 1$	0.5642	4.7711	0.7796	0.2890	0.0481	0.0218	3.4598
$T = 2$	0.6168	4.7430	0.8224	0.2835	0.0517	0.0171	3.5273
$T = 6$	0.6524	2.2821	0.8986	1.0814	0.0729	0.0115	0.9247

Table: Calibrated parameters by expiry for three very short-dated options

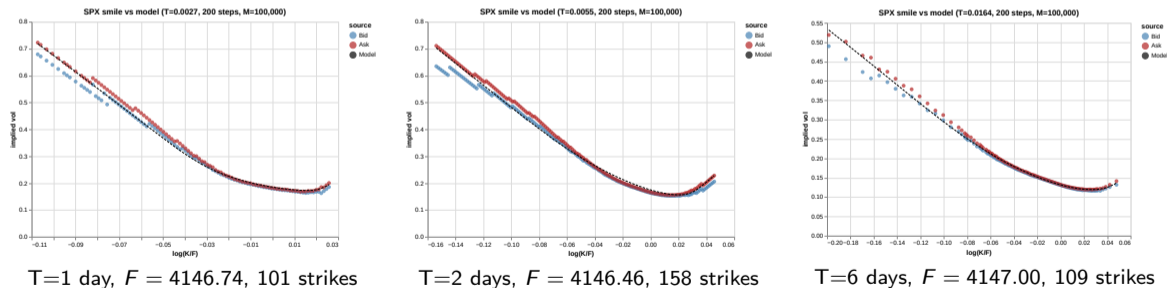


Figure: Market bid-ask implied volatilities and fitted model smile for three very short SPX maturities.

# Model fit for Short-Dated-Options (expiries measured in weeks)

Expiries	$\alpha = H + \frac{1}{2}$	$\lambda$	$\rho$	$c$	$m_0$	$\kappa_0$	$\kappa_2$
$T = 1$	0.6138	3.2444	0.6877	0.7978	0.0747	0.0094	1.2535
$T = 2$	0.5561	2.9191	0.5966	0.2173	0.0634	0.0031	4.6030
$T = 3$	0.6144	1.7727	0.5812	0.2911	0.0864	0.0011	3.4354

Table: Calibrated parameters by expiry for three short-dated-options.

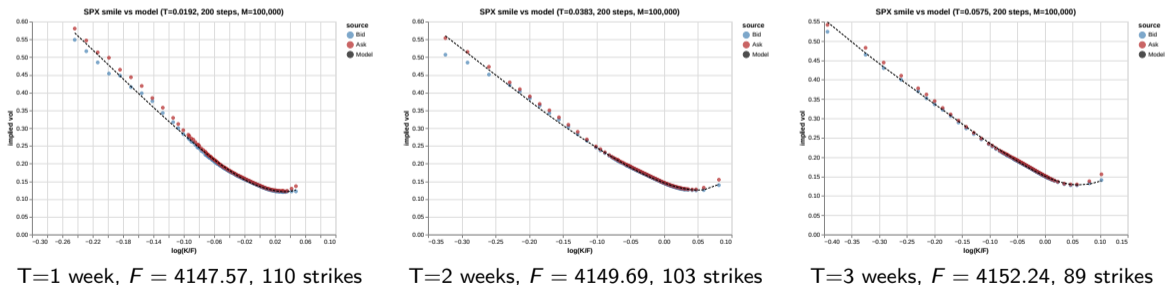
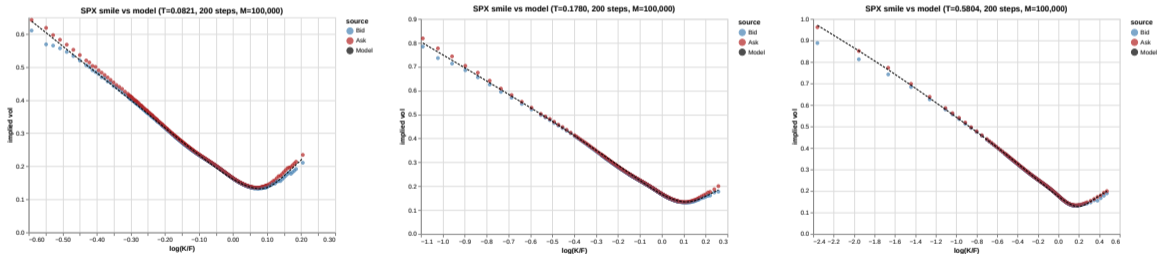


Figure: Market bid–ask implied volatilities and fitted model smile for three short SPX maturities.

# Model fit for Medium-Term Options (expiries measured in months)

Expiries	$\alpha = H + \frac{1}{2}$	$\lambda$	$\rho$	$c$	$m_0$	$\kappa_0$	$\kappa_2$
$T = 1$	0.6386	1.4906	0.8717	0.5330	0.0921	0.0108	1.8762
$T = 2.17$	0.5929	1.3453	0.8985	0.5784	0.1139	0.0077	1.7289
$T = 6.1$	0.5894	1.6437	0.8909	0.2115	0.1002	0.0026	4.7277

Table: Calibrated parameters for three medium term expiries .



$T = 1$ ,  $F = 4155.67$ , 299 strikes

$T=2.17$ ,  $F = 4170.88$ , 280 strikes

$T = 6.1$ ,  $F = 4233.80$  134 strikes

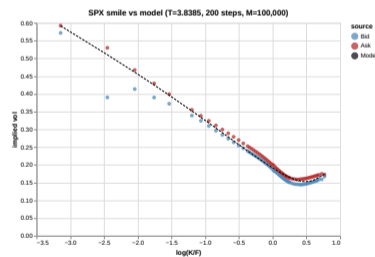
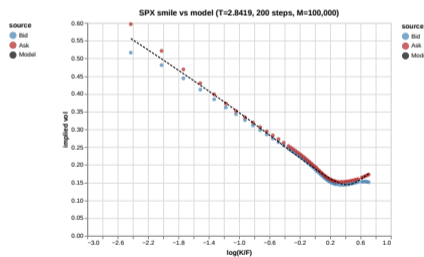
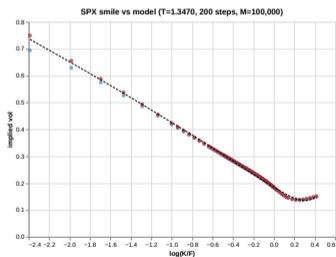
Figure: Market bid-ask implied volatilities and fitted model smile for three medium term SPX maturities.

- The fits are **almost perfect** on the three expiries,
- the model prices **lies between the bid-ask spreads** despite market tightness.

# Model fit for Long-Dated Options (expiries measured in years)

Expiries	$\alpha = H + \frac{1}{2}$	$\lambda$	$\rho$	$c$	$m_0$	$\kappa_0$	$\kappa_2$
$T = 1.35$	0.5838	0.9219	0.8796	1.7510	0.3166	0.0016	0.5711
$T = 2.842$	0.5656	0.7569	0.8869	0.3611	0.1573	0.0003	2.7695
$T = 3.838$	0.5646	0.7736	0.8444	1.0135	0.2768	0.0003	0.9867

Table: Calibrated parameters for three long-dated expiries.



$T = 1.35$ ,  $F = 4357.90$ , 105 strikes

$T = 2.842$ ,  $F = 4535.03$ , 58 strikes

$T = 3.838$ ,  $F = 4643.29$ , 61 strikes

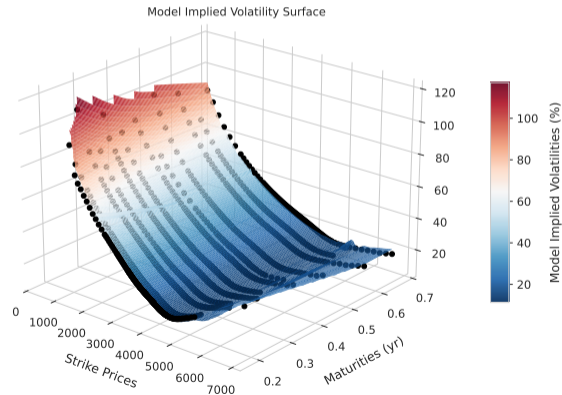
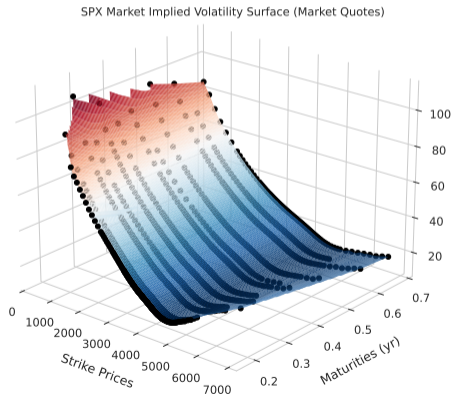
Figure: Market bid-ask implied volatilities and fitted model smile for three long-dated SPX maturities.

- The model provide a **fairly good fit** of the long expiries volatility slices.

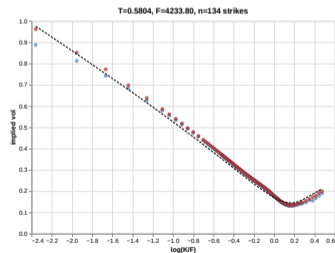
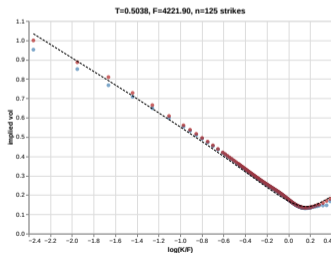
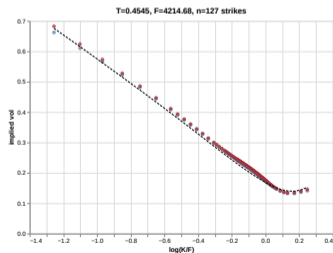
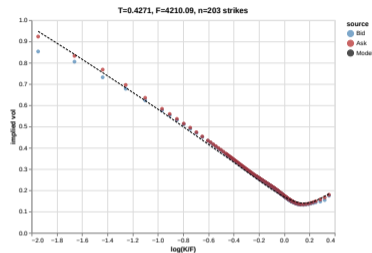
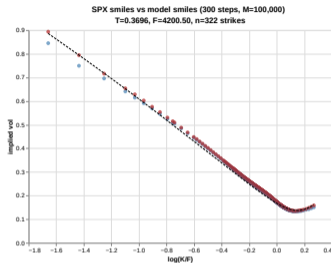
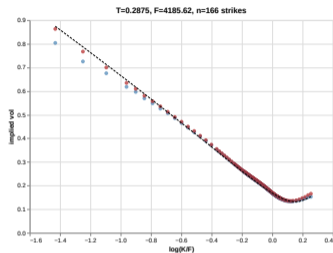
# Joint Calibration of Options: Model-Implied Volatility Surface (From 4 to 7 months)

Expiries (in years)	$\alpha = H + \frac{1}{2}$	$\lambda$	$\rho$	$c$	$m_0$	$\kappa_0$	$\kappa_2$
0.2875, 0.3696, 0.4271, 0.4545, 0.5038, 0.5804	0.5929	1.3453	0.8985	0.5784	0.1139	0.0077	1.7289

Table: Calibrated parameters for the volatility surface.



# Joint Calibration of Options: Model-Implied Volatility Surface



- The model adapt to each patterns of the different expiries, producing a **fairly good fit to the market data.**

# Joint Calibration of Options: Model-Implied Volatility Surface (out of samples), from 1 to 3 weeks

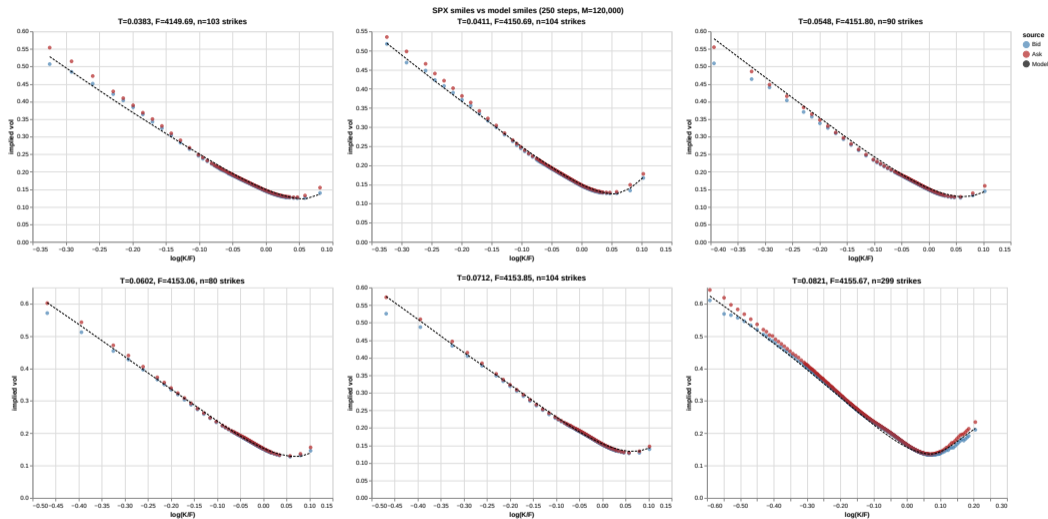


Figure: Calibrated SPX Implied Volatility surface (Volatility smiles for maturities between one week and one month)



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Thanks For Your Attention!

Questions ?